

# 36-763 Hierarchical Linear Models: Homework 5

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## Problem 1

a.

```
> lm1a.1 = lm(Classical ~ 1 + Instrument + Harmony + Voice)
> lm1a.2 = lm(Classical ~ 1 + Harmony + Voice)
> lm1a.3 = lm(Classical ~ 1 + Instrument + Voice)
> lm1a.4 = lm(Classical ~ 1 + Instrument + Harmony)
> # Test fixed effect of Instrument
> anova(lm1a.1, lm1a.2)
```

Analysis of Variance Table

Model 1: Classical ~ 1 + Instrument + Harmony + Voice

Model 2: Classical ~ 1 + Harmony + Voice

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	2485	13108				
2	2487	17235	-2	-4127.6	391.26	< 2.2e-16 ***

---

Signif. codes: 0 \*\*\* 0.001 \*\* 0.01 \* 0.05 . 0.1 1

```
> # Test fixed effect of Harmony
```

```
> anova(lm1a.1, lm1a.3)
```

Analysis of Variance Table

Model 1: Classical ~ 1 + Instrument + Harmony + Voice

Model 2: Classical ~ 1 + Instrument + Voice

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
1	2485	13108				
2	2488	13381	-3	-273.65	17.293	4.107e-11 ***

---

Signif. codes: 0 \*\*\* 0.001 \*\* 0.01 \* 0.05 . 0.1 1

```
> # Test fixed effect of voice
```

```
> anova(lm1a.1, lm1a.4)
```

Analysis of Variance Table

Model 1: Classical ~ 1 + Instrument + Harmony + Voice

Model 2: Classical ~ 1 + Instrument + Harmony

	Res.Df	RSS	Df	Sum of Sq	F	Pr(>F)
--	--------	-----	----	-----------	---	--------

```

1  2485 13108
2  2487 13193 -2    -85.64 8.1181 0.0003061 ***
---
Signif. codes:  0 *** 0.001 ** 0.01 * 0.05 . 0.1 1

```

The three F-tests for the significant of instrument, harmony and voice have p-values of 0, 0 and 0.0003 respectively. We therefore conclude that all three variables are highly significant predictors of Classical ratings.

b.

I

$$Classical_j = \beta_{0[j]} + \beta_1 Instrument + \beta_2 Harmony + \beta_3 Voice + \epsilon_j \stackrel{i.i.d}{\sim} N(0, \sigma^2)$$

$$\beta_{0[j]} = \alpha_0[j] + \epsilon_j \stackrel{i.i.d}{\sim} N(0, \tau^2)$$

II

```

> library(arm)
> library(RLRSim)
> # Define a L-Calc Function that calculates BIC, AIC, DIC, calculated using
> # BIC(), AIC() and display() for DIC
> LCalc <- function(arg1, arg2, n = nargs(), label = c("Model 1", "Model 2")){
+   if (n == 1){
+     mat = cbind(AIC(arg1)[1], BIC(arg1)[1], as.numeric(logLik(arg1)[1]))
+   }
+   if (n == 2){
+     mat = rbind(cbind(AIC(arg1)[1], BIC(arg1)[1], as.numeric(logLik(arg1)[1])),
+       cbind(AIC(arg2)[1], BIC(arg2)[1], as.numeric(logLik(arg2)[1])))
+   }
+   colnames(mat) = c("AIC", "BIC", "DIC")
+   if (n == 1)
+     rownames(mat) = c(paste(label[1]))
+   if (n == 2)
+     rownames(mat) = c(paste(label[1]), paste(label[2]))
+   return(mat)
+ }
> # Using BIC/AIC/DIC
> lmer1b.1 = lmer(Classical ~ (1|Subject) + Instrument + Harmony + Voice)
> lm1b.1 = lm(Classical ~ 1 + Instrument + Harmony + Voice)
> LCalc(lmer1b.1, lm1b.1, n = 2, label = c("Random Effect Intercept", "Fixed Effect Intercept"))

```

	AIC	BIC	DIC
Random Effect Intercept	10491.51	10549.73	-5235.757
Fixed Effect Intercept	11230.45	11282.84	-5606.225

The differences in AIC and BIC are 500-600 lower in the random intercept model compared to the fixed effect intercept model. Based on the rule of thumb that a AIC/BIC difference of 3 is significant, the random effect model is a significantly better fit.

```

> # Using simulated restricted likelihood ratio test
> exactRLRT(lmer1b.1)

```

simulated finite sample distribution of RLRT.

(p-value based on 10000 simulated values)

data:

RLRT = 763.3759, p-value < 2.2e-16

### III

The restricted likelihood ratio test for linear mixed models tests whether the variance of a random effect is 0 in a linear mixed model. Because p-value < 0.001, we reject the null and conclude that the random effect is significant.

```
> lmer1b.2 = lmer(Classical ~ (1|Subject) + Instrument + Harmony + Voice)
> lmer1b.3 = lmer(Classical ~ (1|Subject) + Harmony + Voice)
> lmer1b.4 = lmer(Classical ~ (1|Subject) + Instrument + Voice)
> lmer1b.5 = lmer(Classical ~ (1|Subject) + Instrument + Harmony)
> anova(lmer1b.2, lmer1b.3)
```

Data:

Models:

lmer1b.3: Classical ~ (1 | Subject) + Harmony + Voice

lmer1b.2: Classical ~ (1 | Subject) + Instrument + Harmony + Voice

	Df	AIC	BIC	logLik	deviance	Chisq	Chi	Df	Pr(>Chisq)
lmer1b.3	8	11408	11455	-5696.2	11392				
lmer1b.2	10	10469	10527	-5224.4	10449	943.59		2	< 2.2e-16 ***

---

Signif. codes: 0 \*\*\* 0.001 \*\* 0.01 \* 0.05 . 0.1 1

```
> anova(lmer1b.2, lmer1b.4)
```

Data:

Models:

lmer1b.4: Classical ~ (1 | Subject) + Instrument + Voice

lmer1b.2: Classical ~ (1 | Subject) + Instrument + Harmony + Voice

	Df	AIC	BIC	logLik	deviance	Chisq	Chi	Df	Pr(>Chisq)
lmer1b.4	7	10539	10580	-5262.4	10525				
lmer1b.2	10	10469	10527	-5224.4	10449	75.931		3	2.288e-16 ***

---

Signif. codes: 0 \*\*\* 0.001 \*\* 0.01 \* 0.05 . 0.1 1

```
> anova(lmer1b.2, lmer1b.5)
```

Data:

Models:

lmer1b.5: Classical ~ (1 | Subject) + Instrument + Harmony

lmer1b.2: Classical ~ (1 | Subject) + Instrument + Harmony + Voice

	Df	AIC	BIC	logLik	deviance	Chisq	Chi	Df	Pr(>Chisq)
lmer1b.5	8	10489	10536	-5236.6	10473				
lmer1b.2	10	10469	10527	-5224.4	10449	24.24		2	5.45e-06 ***

---

Signif. codes: 0 \*\*\* 0.001 \*\* 0.01 \* 0.05 . 0.1 1

The effect of the three predictor variables are once again highly significant (p-values = 0) even after the inclusion of the random effect intercepts.

## Part C

### I

```
> lmer9 = lmer(Classical ~ (1|Subject:Instrument) + (1|Subject:Harmony) + (1|Subject:Voice))
> LCalc(lmer9, lm1a.1, n = 2, label = c("Model 1c", "Model 1a"))
```

	AIC	BIC	DIC
Model 1c	10222.76	10251.87	-5106.382
Model 1a	11230.45	11282.84	-5606.225

```
> LCalc(lmer9, lmer1b.1, n = 2, label = c("Model 1c", "Model 1b"))
```

	AIC	BIC	DIC
Model 1c	10222.76	10251.87	-5106.382
Model 1b	10491.51	10549.73	-5235.757

Based on the AIC, BIC and DIC values. The model in 1C with the new random effect terms are significantly better than the all fixed effect model in 1a and the random intercept model in 1b.

### II

```
> summary(lmer9)$varcor
```

Groups	Name	Std.Dev.
Subject:Harmony	(Intercept)	0.73500
Subject:Voice	(Intercept)	0.25639
Subject:Instrument	(Intercept)	1.98726
Residual		1.56497

The estimated variance components are equal to 0.54022500, 0.06573583, 3.94920231, 2.44913110 for Harmony, Voice, Instrument and Residuals respectively. The variance component of the residual is similar in magnitude to the variance component for instrument. It is also marginally larger than the variance component for the harmony. This indicates that the model is appropriate with respect to instrument and harmony. However, the it is significantly larger than the variance component for voice, which indicates that adding a random effect term of the form person:voice might not be appropriate.

### III

$$\begin{aligned} Classical_j &= X_j\beta + \alpha_{ij} + \gamma_{ik} + \delta_{in} + \epsilon_j \stackrel{i.i.d}{\sim} N(0, \sigma^2) \\ \alpha_{ij} &= \psi_{ij} + \eta \stackrel{i.i.d}{\sim} N(0, \kappa^2) \end{aligned}$$

Where  $\alpha_{ij}$  represents the estimate for ith individual, jth Instrument

$$\gamma_{ik} = \theta_{ik} + \nu \stackrel{i.i.d}{\sim} N(0, \lambda^2)$$

Where  $\theta_{ik}$  represents the estimate for ith individual, kth Harmony

$$\delta_{in} = \tau_{in} + \pi \stackrel{i.i.d}{\sim} N(0, \mu^2)$$

Where  $\tau_{in}$  represents the estimate for ith individual, nth Voice

## Problem 2

### Part a

```
> # Fit new models with restricted dataset with all missing values ommited.
>
> lmer10 = lmer(Classical ~ (1|Subject:Instrument) + (1|Subject:Harmony) +
+               (1|Subject:Voice),data = na.omit(ratings))
> sort(apply(apply(ratings, 2, (is.na)),2,mean))
```

	X	Subject
	0.00000000	0.00000000
Harmony		Instrument
	0.00000000	0.00000000
Voice		Selfdeclare
	0.00000000	0.00000000
OMSI		X16.minus.17
	0.00000000	0.00000000
ConsInstr		Instr.minus.Notes
	0.00000000	0.00000000
PianoPlay		GuitarPlay
	0.00000000	0.00000000
first12		Classical
	0.00000000	0.01071429
Popular		ClsListen
	0.01071429	0.01428571
PachListen		Composing
	0.02857143	0.02857143
CollegeMusic		X1990s2000s
	0.04285714	0.05714286
KnowRob	X1990s2000s.minus.1960s1970s	
	0.07142857	0.07142857
APTheory		KnowAxis
	0.08571429	0.11428571
NoClass		ConsNotes
	0.11428571	0.14285714
X1stInstr		X2ndInstr
	0.60000000	0.87142857

```
> names(sort(apply(apply(ratings, 2, (is.na)),2,mean)))
```

[1] "X"	"Subject"
[3] "Harmony"	"Instrument"
[5] "Voice"	"Selfdeclare"
[7] "OMSI"	"X16.minus.17"
[9] "ConsInstr"	"Instr.minus.Notes"
[11] "PianoPlay"	"GuitarPlay"
[13] "first12"	"Classical"
[15] "Popular"	"ClsListen"
[17] "PachListen"	"Composing"
[19] "CollegeMusic"	"X1990s2000s"
[21] "KnowRob"	"X1990s2000s.minus.1960s1970s"
[23] "APTheory"	"KnowAxis"

```

[25] "NoClass"                "ConsNotes"
[27] "X1stInstr"              "X2ndInstr"

> # Calculating the BIC by adding one variable at a time
> BICCalc = NULL
> for (i in c(1:6,8:23))
+ {
+ BICCalc[i] = BIC(update(lmer10, paste("."+names(ratings)[i+5])))
+ }
> # Finding the names with the 5 lowest BIC values
> A=cbind(c(6:11,"NA", 13:28),BICCalc)
> print(A)

      BICCalc
[1,] "6"  "801.876178965197"
[2,] "7"  "811.749212337818"
[3,] "8"  "803.466227711475"
[4,] "9"  "801.949630925903"
[5,] "10" "803.132546582875"
[6,] "11" "803.216903508562"
[7,] "NA" NA
[8,] "13" "803.535693054676"
[9,] "14" "803.60843741513"
[10,] "15" "802.418393816524"
[11,] "16" "802.713999598956"
[12,] "17" "803.194466749332"
[13,] "18" "800.724085412866"
[14,] "19" "801.31020517473"
[15,] "20" "799.766877366058"
[16,] "21" "802.232136926553"
[17,] "22" "800.347033888035"
[18,] "23" "802.985753190926"
[19,] "24" "802.894956626512"
[20,] "25" "802.840333117005"
[21,] "26" "802.365327817851"
[22,] "27" "797.448891980899"
[23,] "28" "789.201356299935"

> names(ratings)[as.numeric(A[order(A[,2]),][3:7,1])]

[1] "APTheory"      "PianoPlay"      "CollegeMusic"   "NoClass"        "Selfdeclare"

> # Construct new dataset with these 5 variables and omit any rows with missing values.
> ratings_new = data.frame(APTheory = ratings$APTheory, PianoPlay = ratings$PianoPlay,
+                           CollegeMusic = ratings$CollegeMusic, NoClass = ratings$NoClass,
+                           Selfdeclare = ratings$Selfdeclare, Instrument = ratings$Instrument,
+                           Voice = ratings$Voice, Harmony = ratings$Harmony,
+                           Classical = ratings$Classical, Popular = ratings$Popular,
+                           Subject = ratings$Subject)
> ratings_new = na.omit(ratings_new)
> lmer11a = lmer(Classical ~ (1|Subject:Instrument) + (1|Subject:Harmony) +
+               (1|Subject:Voice),data = ratings_new)
> lmer11b = lmer(Classical ~ (1|Subject:Instrument) + (1|Subject:Harmony) +

```

```
+ (1|Subject:Voice) + APTheory + PianoPlay + CollegeMusic +
+ NoClass + Selfdeclare,data = ratings_new)
> LCalc(lmer11a, lmer11b, n = 2, label = c("Model 2a", "Model 1C "))
```

	AIC	BIC	DIC
Model 2a	8682.915	8711.185	-4336.458
Model 1C	8694.794	8751.334	-4337.397

We elected to select the 5 most impactful variables, as selected by BIC. After subsetting the original data and after excluding all the missing values, our subset still contained 2109 out of the original 2520 observations. The new model with 5 added predictor variables is a significantly better fit according to BIC, AIC, DIC.

## Part B

```
> lmer11 = lmer(Classical ~ (1|Subject:Instrument) + (1|Subject:Harmony) +
+ (1|Subject:Voice) + APTheory + PianoPlay + CollegeMusic +
+ NoClass + Selfdeclare,data = ratings_new)
> lmer12 = lmer(Classical ~ (1|Subject:Harmony) + (1|Subject:Voice)
+ APTheory + PianoPlay + CollegeMusic + NoClass + Selfdeclare, data = ratings_new)
> lmer13 = lmer(Classical ~ (1|Subject:Instrument) + (1|Subject:Voice)
+ APTheory + PianoPlay + CollegeMusic + NoClass + Selfdeclare, data = ratings_new)
> lmer14 = lmer(Classical ~ (1|Subject:Instrument) + (1|Subject:Harmony)
+ APTheory + PianoPlay + CollegeMusic + NoClass + Selfdeclare, data = ratings_new)
> LCalc(lmer11, lmer12, n = 2, label = c("Full Model", "Without 1:Subject:Instrument"))
```

	AIC	BIC	DIC
Full Model	8694.794	8751.334	-4337.397
Without 1:Subject:Instrument	9946.577	9997.462	-4964.288

```
> LCalc(lmer11, lmer13, n = 2, label = c("Full Model", "Without 1:Subject:Harmony"))
```

	AIC	BIC	DIC
Full Model	8694.794	8751.334	-4337.397
Without 1:Subject:Harmony	8830.276	8881.161	-4406.138

```
> LCalc(lmer11, lmer14, n = 2, label = c("Full Model", "Without 1:Subject:Voice"))
```

	AIC	BIC	DIC
Full Model	8694.794	8751.334	-4337.397
Without 1:Subject:Voice	8698.355	8749.241	-4340.178

Based on BIC, AIC and DIC values, which are tens to hundreds lower in the full model compared to the partial models without 1:Subject:Harmony or 1:Subject:Instrument, we can say that the full model is significantly better fitting than the model without 1:Subject:Instrument or 1:Subject:Harmony. However, the reduced model without 1:Subject:Voice has lower BIC and DIC values than the full model. So therefore, there is evidence to suggest that the random effect of 1:Subject:Voice is not needed when 5 of the most impactful predictors have been added.

## Part C

```
> display(lmer14)
```

```
lmer(formula = Classical ~ (1 | Subject:Instrument) + (1 | Subject:Harmony) +
      APTheory + PianoPlay + CollegeMusic + NoClass + Selfdeclare,
      data = ratings_new)
```

	coef.est	coef.se
(Intercept)	5.72	0.55
APTheory	0.62	0.43
PianoPlay	0.08	0.11
CollegeMusic	0.41	0.44
NoClass	-0.05	0.12
Selfdeclare	-0.22	0.16

Error terms:

Groups	Name	Std.Dev.
Subject:Harmony	(Intercept)	0.76
Subject:Instrument	(Intercept)	2.04
Residual		1.59

---

number of obs: 2109, groups: Subject:Harmony, 236; Subject:Instrument, 177  
 AIC = 8698.4, DIC = 8662.7  
 deviance = 8671.5

On average classical music are rated 5.72 points. Individuals who have taken AP Music Theory class gives 0.62 point in score. For every unit increase in piano playing skill, the score increases by 0.08 point. Those who have taken music class in college give 0.40 point in additional score. Those who have not taken music classes in their life tend to 0.05 point lower in score. Those who strongly identify as being a musician give 0.22 point lower in score for every point increase in their self-identification as a musician.

## Problem 3

```
> ratings_new$Ind.Selfdeclare = 0
> ratings_new$Ind.Selfdeclare = (ratings_new$Selfdeclare > 2)*1
> lmer15 = lmer(Classical ~ (1|Subject:Instrument) + (1|Subject:Harmony) +
+             APTheory + PianoPlay + CollegeMusic + NoClass + Ind.Selfdeclare +
+             Ind.Selfdeclare*APTheory + Ind.Selfdeclare*PianoPlay +
+             Ind.Selfdeclare*CollegeMusic + Ind.Selfdeclare*NoClass, data = ratings_new)
> LCalc(lmer14, lmer15, n = 2, label = c("Model Prob 2C", "Model Prob 3"))
```

	AIC	BIC	DIC
Model Prob 2C	8698.355	8749.241	-4340.178
Model Prob 3	8703.278	8776.779	-4338.639

As for the effect of the interaction terms, self declared musicians who have taken AP Music gives a higher score (0.47), plays piano gives a lower score (-0.04\*PianoPlay), have taken college music classes gives a lower score (-0.50) and have not taken music class gives a lower score (-0.15).

The AIC and BIC values were higher for the new model. DIC values are lower. Because BIC and AIC has much harsher penalties for model complexity, the model for problem 3 was penalized for the introduction of many interaction terms. Even with the interaction terms, the new model is not any better at prediction classical scores than the model found in problem 2, part c.



## Problem 4

### Part A

```
> lmer16 = lmer(Popular ~ (1|Subject) + Instrument + Harmony + Voice, data = ratings_new)
> lm4 = lm(Popular ~ 1 + Instrument + Harmony + Voice, data = ratings_new)
> LCalc(lmer14, lm4, n = 2, label = c("Random Effect Intercept", "Fixed Effect Intercept"))
```

	AIC	BIC	DIC
Random Effect Intercept	8698.355	8749.241	-4340.178
Fixed Effect Intercept	9439.576	9490.461	-4710.788

The random intercept model is once again a better fit than the fixed intercept model.

```
> lmer16 = lmer(Popular ~ (1|Subject:Instrument) + (1|Subject:Harmony) +
+               (1|Subject:Voice) + APTheory + PianoPlay + CollegeMusic +
+               NoClass + Selfdeclare, data = ratings_new)
> lmer17 = lmer(Popular ~ (1|Subject:Harmony) + (1|Subject:Voice) +
+               APTheory + PianoPlay + CollegeMusic + NoClass + Selfdeclare, data = ratings_new)
> lmer18 = lmer(Popular ~ (1|Subject:Instrument) + (1|Subject:Voice) +
+               APTheory + PianoPlay + CollegeMusic + NoClass + Selfdeclare, data = ratings_new)
> lmer19 = lmer(Popular ~ (1|Subject:Instrument) + (1|Subject:Harmony) +
+               APTheory + PianoPlay + CollegeMusic + NoClass + Selfdeclare, data = ratings_new)
> LCalc(lmer16, lmer17, n = 2, label = c("Full Model", "Without Instrument"))
```

	AIC	BIC	DIC
Full Model	8657.694	8714.234	-4318.847
Without Instrument	9683.693	9734.579	-4832.847

```
> LCalc(lmer16, lmer18, n = 2, label = c("Full Model", "Without Harmony"))
```

	AIC	BIC	DIC
Full Model	8657.694	8714.234	-4318.847
Without Harmony	8736.812	8787.698	-4359.406

```
> LCalc(lmer16, lmer19, n = 2, label = c("Full Model", "Without Voice"))
```

	AIC	BIC	DIC
Full Model	8657.694	8714.234	-4318.847
Without Voice	8657.096	8707.982	-4319.548

The effect of instrument is once again highly significant. The model without instrument is a significantly worse fit than the full model. The AIC, BIC and DIC value differences on the order of magnitudes of hundreds. The models without Harmony and Voice are still worse fits than the full models, but only marginally. We therefore conclude that all three predictor variables are significant in the prediction of popular music scores.

### Part B

```
> display(lmer16)
```

**must include all of the  
experimental factors, even if not  
significant**

```
lmer(formula = Popular ~ (1 | Subject:Instrument) + (1 | Subject:Harmony) +
      (1 | Subject:Voice) + APTheory + PianoPlay + CollegeMusic +
      NoClass + Selfdeclare, data = ratings_new)
      coef.est coef.se
```

(Intercept)	4.90	0.51
APTheory	0.07	0.40
PianoPlay	0.00	0.10
CollegeMusic	0.03	0.41
NoClass	0.02	0.11
Selfdeclare	0.18	0.15

Error terms:

Groups	Name	Std.Dev.
Subject:Harmony	(Intercept)	0.65
Subject:Voice	(Intercept)	0.19
Subject:Instrument	(Intercept)	1.88
Residual		1.59

---

number of obs: 2109, groups: Subject:Harmony, 236; Subject:Voice, 177; Subject:Instrument, 177  
 AIC = 8657.7, DIC = 8618.1  
 deviance = 8627.9

On average popular music are rated 4.90 points. Individuals who have taken AP Music Theory class gives 0.07 point in score. The effect of piano playing is negligible. Those who have taken music class in college give 0.03 point in additional score. Those who have not taken music classes in their life tend to 0.02 point lower in score. Those who strongly identify as being a musician give 0.18 point higher in score for every point increase in their self-identification as a musician.

## Part C

```
> lmer20 = lmer(Popular ~ (1/Subject:Instrument) + (1/Subject:Harmony) +
+               APTheory + PianoPlay + CollegeMusic + NoClass + Ind.Selfdeclare +
+               Ind.Selfdeclare*APTheory + Ind.Selfdeclare*PianoPlay +
+               Ind.Selfdeclare*CollegeMusic + Ind.Selfdeclare*NoClass, data = ratings_new)
> LCalc(lmer20, lmer16, n = 2, label = c("Model Prob4c", "Model Prob4a"))
```

	AIC	BIC	DIC
Model Prob4c	8658.083	8731.584	-4316.041
Model Prob4a	8657.694	8714.234	-4318.847

The AIC and BIC values were higher for the new model. DIC value was lower because the new model contained more predictor variables, which lead to a lower deviance. The complexity of the new model was heavily penalized by AIC and more so by BIC. Even with the interaction terms, the new model is not any better at prediction popular music scores than the model found in problem 4, part A.

## Problem 5

### Classical and Popular Music Differentiation by Instrument, Harmony and Voice

The classification of aural stimuli as either 'classical' or 'popular' is heavily influenced by the instruments used, the harmonic motion present and the type of voice leading. An anova analysis of a simple linear regression in problem 1, part A indicates that all three of instrument, harmonic motion, and voice leading are significant predictors of both classical and popular music.

#### Repeated Measure Model

By fitting a repeated measures model with a random intercept for each participant, we were able to better estimate the popular and classical scores when compared to a simple linear model. The effects of the instruments used, the harmonic motion present and the type of voice leading was once again significant even after taking into account the personal bias of each of the 36 participants.

#### Random Intercept Model

However, the bias for each of the 36 participants may not be personal, but is rather a function of instrument, harmonic motion or voice leading. After fitting a model with varying intercepts for each of the participant and instrument, harmonic motion and voice leading pairs, the estimated variance components for Subject:Voice and variance components for the residuals indicates that voice is not an appropriate predictor in this model. Therefore, we conclude that the 36 participants have significant biases toward specific types of harmonic motion and instrument, but were much more forgiving in terms of the voice leading.

#### Personal Bias Model

After adding in five of the most significant predictors of classical and popular - whether the participant has taken AP Music, plays piano, has taken college music classes, has ever taken a music class, or self-identifies as a musician. The new model with 5 added predictors is significantly better at predicting popular and classical scores compared to our random intercept model.

#### Musicians vs Non-Musicians Model

When the predictor variable 'SelfIdentify' was dichotomized into roughly two equal levels, the interactions between the indicator for self-identifying musicians and participants who plays the piano, have taken college music classes, or have not taken music classes were significant. Individuals who have more exposure to music are better able to distinguish classical music. Most significantly, participants who have taken AP Music gave classical stimuli a score that is 0.47 higher! However, being a musician only marginally increases one ability to identify popular music stimuli.

not sure what data this is based on. there are no gold standards in this study

#### Conclusion

Instrument and harmony are the two most significant predictors of popular and classical music identification. Participants have significant biases toward the identification of classical and popular music. Some participants are better than other participants with similar qualifications at identifying either popular or classical music. Participants also have significant biases toward different types of instrument and harmony. Being a musician increases one's ability to identify classical music, but only marginally increases one's ability to identify popular music.